

Testing the identification of causal effects in observational data

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In a nutshell

We want to provide a test for identification of a causal effect based on jointly testing for selection-on-observables and instrument validity in observational data.

Identifying assumptions in causal inference

- Causal inference typically makes use of identifying assumptions that are deemed to be untestable.
- The selection-on-observables assumption e.g. imposes that the treatment is as good as randomly assigned after controlling for observed covariates.
- Whether the set of covariates is sufficient is typically motivated by economic theory, intuition, or previous empirical findings.
- However, the plausibility of the selection-on-observables assumption often appears 'shaky' and it would be nice to have a statistical test for the identifying assumptions.

Contribution

- This paper demonstrates that there exists a testable condition for the selection-on-observables assumption in observational data when imposing some causal structure.
- The condition relies on two types of observables, namely covariates to be controlled for and a suspected instrument, and arises if...

Assumption 1 there is no reverse causality from the outcome to the treatment, covariates, or the suspected instrument and from the treatment to the covariates or the suspected instrument,

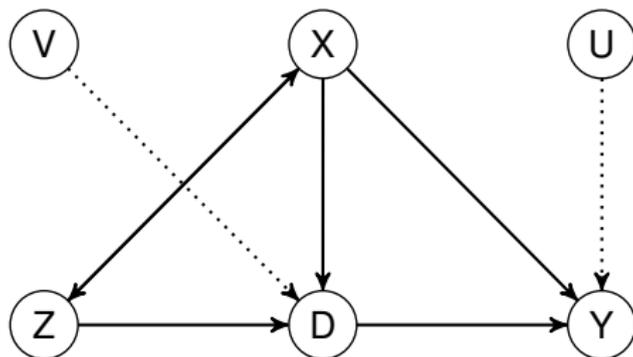
Assumption 2 the suspected instrument and the treatment are conditionally dependent given the covariates.

Notation

- D : Treatment.
- Y : Outcome.
- X : Covariates.
- Z : Suspected instrument.
- V : Unobservables affecting D .
- U : Unobservables affecting Y .
- $Y(d)$: Potential outcome under treatment d .

Graphical illustration

Figure: Causal graph satisfying Assumption 1 and Assumption 2



Conditional independence of the treatment

$$Y(d) \perp\!\!\!\perp D | X \quad \forall d \in \mathcal{D}.$$

- Conditional on X , D is not associated with unobserved characteristics affecting the outcome.

Conditional independence of the instrument

$$Y(d) \perp\!\!\!\perp Z | X \quad \forall d \in \mathcal{D}.$$

- Conditional on X , Z does not directly affect the outcome and is not associated with unobserved characteristics affecting the outcome.

Theorem

Conditional on Assumptions 1 and 2, it holds that

$$Y(d) \perp\!\!\!\perp D|X, \quad Y(d) \perp\!\!\!\perp Z|X \iff Y \perp\!\!\!\perp Z|D = d, X \quad \forall d \in \mathcal{D}.$$

Hence, the testable implication

$$Y \perp\!\!\!\perp Z|D = d, X$$

is necessary and sufficient for the identification of the causal effect when considering potential outcomes $Y(d)$ matching the factual treatment assignment $D = d$.

Testing the identification of average effects

Theorem

Conditional on Assumptions 1 and a (stronger) version of Assumption 2 w.r.t. means, it holds that

$$\begin{aligned} E[Y(d)|D, X] &= E[Y(d)|X], & E[Y(d)|Z, X] &= E[Y(d)|X] \\ \iff E[Y|Z, D = d, X] &= E[Y|D = d, X] & \forall d \in \mathcal{D}. \end{aligned}$$

The testable implication

$$E[Y|Z, D = d, X] = E[Y|D = d, X]$$

is necessary and sufficient for the identification of average effects.

Testing the identification of average effects

Denote $\mu(z, d, x) = E[Y|Z = z, D = d, X = x]$. Considering a binary instrument Z , conditional mean independence is equivalent to:

$$H_0 : \mu(1, d, x) - \mu(0, d, x) = 0 \quad \forall d \in \mathcal{D}, \text{ and } x \in \mathcal{X}.$$

One approach is to test the condition on average (for a binary Z):

$$H_0 : E[\mu(1, D, X) - \mu(0, D, X)] := \Delta = 0.$$

Testing based on DML

For a binary Z , estimate Δ by double machine learning (DML), see e.g. Chernozhukov et al (2018):

$$\Delta = E[\phi(D, X)] \quad (1)$$

with

$$\phi(D, X) = \mu(1, D, X) - \mu(0, D, X) + \frac{(Y - \mu(1, D, X)) \cdot Z}{p(D, X)} - \frac{(Y - \mu(0, D, X)) \cdot (1 - Z)}{1 - p(D, X)}.$$

- $\mu(Z, D, X)$ and p-score $p(D, X)$ are nuisance terms in Neyman-orthogonal score $\phi(D, X)$.
- Under $n^{-1/4}$ -consistency of estimators of $\mu(Z, D, X)$ and $p(D, X)$, the sample analog of (1) is \sqrt{n} -consistent.
- Averaging over X may decrease asymptotic power since violations across X can cancel out.

DML-based testing in subsets

- To increase asymptotic power, test the null in subsamples defined upon values of X, D which are predictive for Δ .
- Procedure:
 - 1) Split sample randomly into two non-overlapping subsamples.
 - 2) In the first subsample, estimate $\phi(D, X)$ and predict it as a function of (D, X) using machine learning.
 - 3) In the second subsample, conduct hypothesis tests using subsets defined as a function of best predictors in (D, X)
- Formally, for L_m denoting the m th subset, one tests

$$H_0 : E[\mu(1, D, X) - \mu(0, D, X) | (D, X) \in L_m] := \Delta(L_m) = 0.$$

- Using multiplier bootstrap, one can construct joint tests across all subsets based on the norm of $\Delta(L_m)$ for $m = 1, \dots, M$.

Testing based on squared differences

Testing for violations across all values of D and X globally, by verifying the following null hypothesis H_0 :

$$E[(\mu(1, D, X) - \mu(0, D, X))^2] := \theta = 0.$$

- Testing is based on the Neyman-orthogonal function $(\mu(1, D, X) - \mu(0, D, X))^2 + \zeta$, where ζ is a mean-zero random variable with bounded variance σ_ζ^2 from below.
- ζ avoids a degenerate variance under H_0 (which would violate asymptotic normality).
- Asymptotic normality holds under H_0 (but not under a violation of H_0).

Application to Angrist and Evans (1998) data

143,410 observations of married white couples with mother's education amounting to 12 years from the 1980 wave of the US Census Public Use Micro Samples.

- Z : sex ratio of a mother's first two children,
- D : 1 for three or more kids, 0 for two kids,
- Y : mother's weeks in employment per year,
- X : mother's age, mother's age at first birth, father's age, and father's income.

Results based on DML and squared differences

Table: Empirical application to the sibling sex instrument

	est	se	pval
DML	0.22	0.11	0.05
Squared difference	0.5	0.02	0.00

Notes: columns 'est', 'se', and 'pval' provide the estimates, the standard errors and the p-values for Δ and θ , respectively.

- The tests point to a violation of conditional independence given our limited set of covariates.

Results based on DML in subsets

Table: Empirical application to the sibling sex instrument

splitting variable	subset	est	se	pval	bonf	boot (l_∞)	boot (l_2)
father's income	in lower half	0.076	0.162	0.641	0.0606	0.0588	0.0825
	in upper half	0.328	0.152	0.030			
treatment D	$D = 1$	0.080	0.175	0.646			
	$D = 0$	0.311	0.144	0.031			
mother's age	in lower half	0.321	0.156	0.039			
	in upper half	0.139	0.159	0.382			
mother's age (afb)	in lower half	0.200	0.168	0.234			
	in upper half	0.226	0.148	0.128			
father's age	in lower half	0.348	0.156	0.026			
	in upper half	0.116	0.159	0.464			
father's income	in lower half	0.076	0.162	0.641	0.2555	0.2225	0.0089
	in upper half	0.328	0.152	0.030			

Notes: columns 'est', 'se', and 'pval' provide the estimates for Δ and the respective standard errors and p-values within subsets. The three last columns 'bonf', 'boot (l_∞)', and 'boot (l_2)' yield the p-value for the joint significance across all subsets using Bonferroni correction and multiplier bootstrap based on the l_∞ -norm and l_2 -norm, respectively.

- Father's income most predictive variable for heterogeneity in violations of conditional independence.
- Joint tests point to violation of conditional independence.

- We suggested a testable condition for the joint satisfaction of selection-on-observables and IV validity assumptions, if the instrument and treatment are conditionally dependent.
- We proposed several machine learning-based tests which permit controlling for covariates in a data-driven way when testing.
- We presented an application to the evaluation of the effects of fertility on female labour supply.

Thank you for your attention!

Our paper is available on arXiv:

<https://arxiv.org/abs/2203.15890>

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